Performance evaluation: Point of view Reliability

System reliability

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Definition

It's the probability of successful operation of a system or system component itself during a given time, reliability is a dimension that is not the equivalent of "quantity", "value" of the system considered. Corresponding to the degree of confidence that can be placed in a machine or mechanism. We note that reliability has become essential since the equipment was complicated

Motivation

Failures in airplanes, rockets or nuclear plants quickly become catastrophic; it is necessary to accurately predict the uptime of each of these systems. Currently, this study is the same time as the project construction

Definition and Notation

Reliability:

- R(t) = Probability (S don't fail on [0,t])
- R(t) is a non increasing function varing between 1 à 0 on [0, +∞ [

Availability:

- Availability A (t) is the probability that the system S is not in default at time t. Note that in the case of non-repairable systems, the definition of A (t) is equivalent to the reliability : A(t) = Probability (S is not default at t)
- Maintenability:
- Maintainability M (t): the probability that the system is repaired on the interval [0 t] knowing that he has failed at time t = 0:
- M(t)=Probability (S is repaired on [0 t]/ S is failed at t=0)
- This concept applies only to repairable systems
- M(t) is a non decreasing function varying between 0 à 1 on [0, +∞ [

Definitions et notations

- Mean time before failures:
- The average duration of system work time before the first failure : «

 Mean Time To Failure »

$$MTTF = \int_{0}^{+\infty} R(t) dt = \int_{0}^{+\infty} tf(t) dt$$

- Mean time to repair:
- The average duration of reparation action: « Mean Time To Repair»

$$MTTR = \int_{0}^{+\infty} (1 - M(t)) dt$$

Definitions et notations

Mean up time

- MUT:« Mean Up Time». It is different to MTTF because when the system is returned to service after a failure, all breakdown elements have not necessarily been repaired
 - Mean down time:
 - MDT:« Mean Down Time». This average corresponds to the detection of the failure, duration of intervention, the duration of the repair and the ready time
 - Mean time between failure:
- MTBF:« Mean Time Between Failure». Mean time between successive failures
 - MTBF=MUT +MDT

MTTF≅MUT

stochastic Processes

Renewal process:

We consider a set of elements whose life is a continuous random variable F with a probability density f. At time t = 0 is put into service the first element and replaced by the following when a failure at time F1. If Fr is the life of the r-th service element, its failure will occurs at date kr, defined by: kr = F1 + F2 +..... Fr

We called renewal function the average value of the number of rotation N (t) occurring on (0, t), the introduction of the first element at time t = 0 is not counted as a renewal. H (t) = E [N (t)]

Called renewal density h (t) derivative H (t).

stochastic Processes

- We called variable renewal process a renewal process for which the random variable F1 has a different density than other random variables Fi.
- We Called residual life Vt the random variable representing the remaining life of the item in service at time t

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Fondamental relations

 We note by T the continuous random variable characterizing the up time of the system

Re liabilityé:
$$R(t) = P(T > t)$$

Failure fuction T: $F(t) = P(T \le t) = 1 - R(t)$
 $F(t)$ failure probability on [0 t]

failure density : $f(t) = \frac{dF(t)}{dt} = -\frac{dR(t)}{dt}$

Le temps moyen de bon fonctionnement:

MTTF = $\int_{0}^{+\infty} tf(t) dt = \int_{0}^{+\infty} R(t) dt$

Mean time to repair : MTTR =
$$\int_{0}^{+\infty} \left[1 - M(t) \right] dt$$

Relations fondamentales

Failure rate and repair rate

Failure rate
$$\lambda(t) = \frac{-\frac{dR(t)}{dt}}{R(t)} \Rightarrow R(t) = \exp\left(-\int_{0}^{t} \lambda(u) du\right)$$

repaire rate : $\mu(t) = \frac{\frac{dM(t)}{dt}}{1 - M(t)} \Rightarrow M(t) = 1 - \exp\left(-\int_{0}^{t} \mu(u) du\right)$
failure density : $f(t) = -\frac{dR(t)}{dt} = \lambda(t)R(t) = \lambda(t)\exp\left(-\int_{0}^{t} \lambda(u) du\right)$
Repair density : $m(t) = \frac{dM(t)}{dt} = \mu(t)[1 - M(t)] = \mu(t)\exp\left(-\int_{0}^{t} \mu(u) du\right)$

Experimentation

 The Principe consists at making N new materials working at t=0 assuring the same working conditions.

- Case 1 N≥50 : Estimation by interval
 - Note the failure date of every material
 - Note the minimal failure date tmin
 - Note the maximal failure date tmax
 - Calculate class number nc= √N (square root on N)
 - calculate the class length Lc=(tmax-tmin)/nc
 - Calculate ni; the number of material failed inside the class i
 i∈{1,....nc}
 - Calculate nsi, the number of surviving material at the beginning of every class i

Method of determination of the material failure law

« New material »

- Case 1 N≥50 : Estimation by interval
 - Estimation of a failure law for every class
 - *probability density function for class i:

* Failure rate for class i:

* Reliability for class i

 * probability distribution function associated with the time to failure for class i

Method of determination of the material failure law

« New material »

- Case 1 N≥50 : Estimation by interval
 - We plot the curve of Ri according to class i (histogram)
 - Using mathematical Software in order to smooth the curve and determine the mathematical expression of R(t)

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(LABFIT, STATFIT...)
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Then we can deduce all the expressions $F(t),f(t),\lambda(t),MUT$ Using theses expression in order to propose :

- An optimal warranty period
- An optimal maintenance plan

-

Application : industrial example (N≥50)

- Case 2 N<50 : Punctual Estimation
 - Note the failure date of every material
 - classify the failure date by increasing order (t₁,t₂,.....t_N)
 - Let "i" representing the failure date order
 - For 20<N<50 (estimation by "rang moyen")
 - probability distribution function associated with the time to failure according to t_i:

$$Fi=i/(N+1)$$

Case 2 N<50 : Punctual Estimation

- For N<20 (estimation by "rang median")
- probability distribution function associated with the time to failure according to t_i:

$$Fi=(i-0.3)/(N+0.4)$$

- Plote Fi according to t_i
- Using mathematical Software in order to smooth the curve and determine the mathematical expression of F(t)

(LABFIT, STATFIT...)

- Then we can deduce all the expressions $R(t),f(t),\lambda(t),MUT$ Using theses expression in order to propose :
- An optimal warranty period
- An optimal maintenance plan

-

Application: industrial example (N<50)

Acceptance test for obtained law Case 1 N≥50 : KHI-Deux Test

- Compute E:
- $E = \sum ((ni-N*Pi)^2)/(N*Pi)$
- And Pi= R(t_{i-1})-R(t_i) with t_{i-1} and t_i are respectively the born inf and sup of every interval I

R is law obtained from the mathematical Software

- γ= nc-k-1 (k the number of parameters of the considered law
- α the value of the risk proposed by the industrial
- Note the value of χ (γ , α) in the Khi-Deux table
- If E> χ (γ , α) the law proposed is rejected
- If E≤ χ (γ, α) the law proposed is accepted

If the law is rejected we move to test another law

Acceptance test for obtained law

- Case 2 N<50 : Klomorgov-Smirnov Test
- Compute D⁺ and D⁻
- $D^+ = \max \{(i/N) F(t_i)\}$, and $D^- = \max\{F(t_i) ((i-1)/N)\}$ ($\forall i \in \{1, 2, ..., N\}$)
- F is law obtained from the mathematical Software
- Compute D= max (D^{+,} D⁻)
- α the value of the risk proposed by the industrial
- Note the value of D_{αN} in the Klomorgov-Smirnov Table
- If D> D_{α,N} the law proposed is rejected
- If D≤ D_{α,N} the law proposed is accepted

Principal law used in industry and research in reliability frame

Usuel discret law

Dirac:

It's a constant law

$$X(\omega) = a, \forall \omega \in \Omega$$

$$F(x) = \begin{cases} 0 & \text{si } x \le a \\ 1 & \text{si } x > a \end{cases}$$
$$E(X) = a$$
$$V(X) = 0$$

Bernoulli:

$$X(\omega) = 1_{A}(\omega) = \begin{cases} 1 & \text{si } \omega \in A \\ 0 & \text{si } \omega \in \overline{A} \end{cases}$$

Parameter is p defined by p=P(A), notation $X \rightarrow B(1,p)$

$$X(\omega) = egin{cases} 1 & p \ 0 & \mathbf{q} = 1 \text{-p} \end{cases}$$
 $F(x) = egin{cases} 0 & \sin x \leq 0 \ \mathbf{q} & \sin 0 < x \leq 1 \ 1 & \sin x > 1 \end{cases}$
 $E(X) = p$
 $V(X) = pq = p(1-p)$

Dem FIGURE EXEMPLE page 66 67

w binomiale »:Parameters *n* and *p*=P(A)

$$\begin{aligned} & P_X(X=k) = C_n^k p^k (1-p)^{n-k} \\ & E(X) = np \\ & V(X) = npq \end{aligned}$$

Notation $X \rightarrow B(n,p)$

« Poisson » :

Parameters λ>0

$$P_{X}(X = k) = e^{-\lambda} \frac{\lambda^{k}}{k!}$$

$$E(X) = \lambda$$

$$V(X) = \lambda$$

Notation $X \rightarrow P(\lambda)$

« Pascal »:

Parameter k

$$A \longrightarrow A \longrightarrow A$$

Si on pose p = P(A) la probabilité de cette évènement est:

$$P_k(X = k) = (1-p)^{k-1} p$$

$$E(X) = \frac{1}{p}$$

$$V(X) = \frac{q}{p^2}$$

$$V(X) = \frac{q}{p^2}$$

« binomiale négative »: Parameters *n* and *y*

On a n - 1 réalisations de l'évènement A au cours de y - 1 premières épreuves et qui se conclut par l'évènement A. On déduit la probabilité individuelle :

$$P_k(Y = y) = C_{y-1}^{n-1}(1-p)^{y-n}p^n$$

$$E(Y) = \frac{n}{p}$$

$$E(Y) = \frac{n}{p}$$

$$V(Y) = \frac{nq}{p^2}$$

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Continuous law

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« Loi uniforme »

$$f(x) = \begin{cases} k & \text{si } x \in [a,b] \\ 0 & \text{si non} \end{cases}$$

La densité
$$f(x) = \frac{1}{b-a} \forall x \in [a,b]$$

La fonction de répartition: $F(x) = \begin{cases} 0 & \text{si } x < a \\ \frac{x-a}{b-a} & \text{si } a \leq x < b \\ 1 & \text{si } b \leq x \end{cases}$

L'espérance: $E(x) = \frac{a+b}{2}$
 $V(x) = \frac{(b-a)^2}{12}$

En particulier a=0 et b=1

La densité $f(x) = \begin{cases} 1 & \forall x \in [0,1] \\ 0 & \text{sinon} \end{cases}$

La fonction de répartition: $F(x) = \begin{cases} 0 & \text{si } x < 0 \\ x & \text{si } 0 \leq x < 1 \\ 1 & \text{si } 1 \leq x \end{cases}$

L'espérance: $E(x) = \frac{1}{2}$

La variance: $V(x) = \frac{1}{12}$

Exponential law :

$$f(x) = \begin{cases} \theta e^{-\theta x} & \text{si } 0 \le x \\ 0 & \text{si } x < 0 \end{cases}$$

La fonction de répartition : $F(x) = 1 - e^{-\theta x}$

L'espérance:
$$E(x) = \frac{1}{\theta}$$

La variance :V(x) =
$$\frac{1}{\theta^2}$$

Notation $X \rightarrow \varepsilon(\theta)$

- Laplace-Gauss:
- Parameters m and σ

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp{-\left(\frac{(x-m)^2}{2\sigma^2}\right)}$$

L'espérance: E(x) = m

La variance :V(x) = σ^2

Notation $X \rightarrow N(m, \sigma)$

« gamma »

Parameters p>0 and θ >0

$$f(x) = \frac{\theta^{p}}{\Gamma(p)} e^{-(\theta x)} x^{p-1}, x \ge 0$$

$$A \text{vec:} \Gamma(p) = \int_{0}^{+\infty} e^{-(x)} x^{p-1} dx$$

$$L'\text{espérance:} E(x) = \frac{p}{\theta}$$

$$La \text{ variance:} V(x) = \frac{p}{\theta^{2}}$$

Lois usuelles continues

« Khi-Deux »:

Gamma with p=n/2 and θ =1/2 (γ (n/2, 1/2))

$$f(x) = \frac{1}{2^{\left(\frac{n}{2}\right)}\Gamma\left(\frac{n}{2}\right)} e^{-\left(\frac{x}{2}\right)} x^{\frac{n}{2}-1}, x \ge 0$$
L'espérance: $E\left(\chi_n^2\right) = \frac{n/2}{1/2} = n$
La variance : $V\left(\chi_n^2\right) = \frac{n/2}{1/4} = 2n$

« Beta":

Second:

Si X = $\gamma(p)$ and Y= $\gamma(q)$, we deduce Z=X/Y = $\beta_{11}(p,q)$

$$f(z) = \frac{1}{\beta(p,q)} \frac{z^{p-1}}{(1+z)^{p+q}}, z \ge 0$$

$$avec : \beta(p,q) = \frac{\Gamma(p)\Gamma(q)}{\Gamma(p+q)}$$
L'espérance: $E(Z) = \frac{p}{q-1} \neq 0$
La variance : $V(Z) = \frac{p(p+q-1)}{(q-1)^2(q-2)}$

$$B(x,y) = \int_0^1 t^{x-1} (1-t)^{y-1} dt$$

- « Beta »:
- First

$$T = \frac{X}{X+Y} = \frac{Z}{1+Z}$$
Sa densité pour $0 \le t \le 1$

$$f(t) = \frac{1}{\beta(p,q)} t^{p-1} (1-t)^{q-1}$$
L'espérance: $E(T) = \frac{p}{p+q}$
La variance : $V(T) = \frac{pq}{(p+q)^2(p+q+1)}$

« log-normale »:

Parameters m and σ

La fonction de répartition:
$$F(x) = \phi\left(\frac{\ln x - m}{\sigma}\right)$$

La densité est: $f(x) = \frac{1}{\sigma x} \phi\left(\frac{\ln x - m}{\sigma}\right) = \frac{1}{\sigma x \sqrt{2\pi}} \exp\left(-\frac{1}{2\sigma^2} (\ln x - m)^2\right)$
 $AVEC: \phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{x} e^{-\left(\frac{u^2}{2}\right)}$
 $E(X) = \exp\left(m + \frac{\sigma^2}{2}\right)$
 $V(X) = \exp\left(2m + \sigma^2\right) \left(\exp\left(\sigma^2\right) - 1\right)$

« Pareto »:

Parameters x_0 ($x \ge x_0 > 0$) and $\alpha > 0$:

La densité est:
$$f(x) = \frac{\alpha}{x_0} \left(\frac{x_0}{x}\right)^{\alpha+1}$$

$$E(X) = \frac{\alpha x_0}{\alpha - 1} \text{ avec } \alpha > 1$$

$$V(X) = \frac{\alpha x_0^2}{(\alpha - 1)^2 (\alpha - 2)} \text{ avec } \alpha > 2$$

Lois Weibull trois paramètres

Densité de probabilité :

$$f(x; \beta, \lambda, \theta) = \left(\frac{\beta}{\lambda}\right) \left(\frac{x - \theta}{\lambda}\right)^{(\beta - 1)} e^{-\left(\frac{x - \theta}{\lambda}\right)^{\beta}}$$

Fonction de répartition :

$$F(x;\beta,\lambda,\theta) = 1 - e^{-\left(\frac{x-\theta}{\lambda}\right)^{\beta}}$$

Lois Weibull deux paramètres (B,A)

Densité de probabilité :

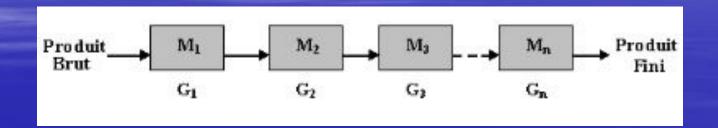
$$f(x; \beta, \lambda) = \left(\frac{\beta}{\lambda}\right) \left(\frac{x}{\lambda}\right)^{(\beta-1)} e^{-\left(\frac{x}{\lambda}\right)^{\beta}}$$

Fonction de répartition :

$$F(x;\beta,\lambda) = 1 - e^{-\left(\frac{x}{\lambda}\right)^{\beta}}$$

Structures

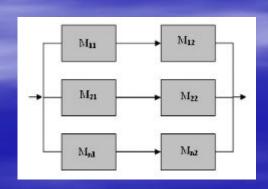
series

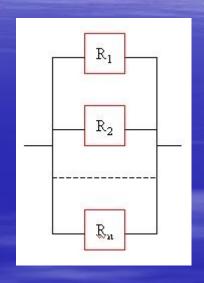


Structures

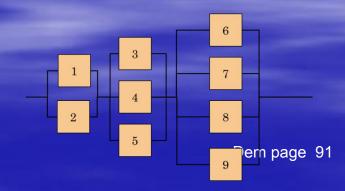
parallel

Series-parallel





Parallel-series



- Complex Structures
- Bridge system

Theorem of Bays

Exampl

Structures

- series
- parallel
- Parallel-series
 - Series-parallel

Structures

- series
- parallel
- Parallel-series
 - Series-parallel

Thank you for attention